$$\bar{\lambda}_{i} = \lambda_{i} \qquad \text{for } i < k$$
$$\tilde{\lambda}_{k} = \frac{(1 - \rho_{1} - \dots - \rho_{k-1})}{\bar{X}_{k}}$$

For priorities i < k the arrival process is Poisson so the same calculation for the waiting time as before gives

$$W_{i} = \frac{\sum_{i=1}^{k} \tilde{\lambda}_{i} \overline{X}_{i}^{2}}{2(1 - \rho_{1} - \dots - \rho_{i-1})(1 - \rho_{1} - \dots - \rho_{i})}, \qquad i < k$$

For priority k and above we have infinite average waiting time in queue.

3.40

(a) The algebraic verification using Eq. (3.79) listed below

$$W_k = R/(1 - \rho_1 - \ldots - \rho_{k-1})(1 - \rho_1 - \ldots - \rho_k)$$

is straightforward. In particular by induction we show that

$$\rho_1 W_1 + \dots + \rho_k W_k = \frac{R(\rho_1 + \dots + \rho_k)}{1 - \rho_1 - \dots - \rho_k}$$

The induction step is carried out by verifying the identity

$$\rho_1 W_1 + \dots + \rho_k W_k + \rho_{k+1} W_{k+1} = \frac{R(\rho_1 + \dots + \rho_k)}{1 - \rho_1 - \dots - \rho_k} + \frac{\rho_{k+1} R}{(1 - \rho_1 - \dots - \rho_k)(1 - \rho_1 - \dots - \rho_{k+1})}$$

The alternate argument suggested in the hint is straightforward.

(b) Cost

$$C = \sum_{k=1}^{n} c_k N_Q^k = \sum_{k=1}^{n} c_k \lambda_k W_k = \sum_{k=1}^{n} \left( \frac{c_k}{X_k} \right) \rho_k W_k$$

We know that  $W_1 \le W_2 \le \dots \le W_n$ . Now exchange the priority of two neighboring classes i and j=i+1 and compare C with the new cost

$$C' = \sum_{k=1}^{n} \left( \frac{c_k}{\overline{X}_k} \right) \rho_k W'_k$$

In C' all the terms except k = i and j will be the same as in C because the interchange does not affect the waiting time for other priority class customers. Therefore

$$C'-C = \frac{c_j}{\overline{X_j}} \rho_j W'_j + \frac{c_i}{\overline{X_i}} \rho_i W'_i - \frac{c_i}{\overline{X_i}} \rho_i W_i - \frac{c_j}{\overline{X_i}} \rho_j W_j.$$

We know from part (a) that

$$\sum_{k=1}^{n} \rho_{k} W_{k} = \text{constant.}$$

Since  $W_k$  is unchanged for all k except k = i and j (= i+1) we have

$$\rho_i W_i + \rho_j W_j = \rho_i W'_i + \rho_j W'_j.$$

Denote

$$B = \rho_i W'_i - \rho_i W_i = \rho_j W_j - \rho_j W'_j$$

Clearly we have  $B \ge 0$  since the average waiting time of customer class i will be increased if class i is given lower priority. Now let us assume that

$$\frac{c_i}{\overline{X}_i} \le \frac{c_j}{\overline{X}_j}$$

Then

$$C'-C = \frac{c_i}{\overline{X_i}} \left( \rho_i W'_i \rho_i W_i \right) - \frac{c_j}{\overline{X_j}} \left( \rho_j W_j \rho_j W'_j \right) = B\left( \frac{c_i}{\overline{X_i}} - \frac{c_j}{\overline{X_j}} \right)$$

Therefore, only if  $\frac{c_i}{\overline{X}_i} < \frac{c_{i+1}}{\overline{X}_{i+1}}$  can we reduce the cost by exchanging the priority order of i and i+1. Thus, if (1,2,3,...,n) is an optimal order we must have

$$\frac{c_1}{\overline{X_1}} \ge \frac{c_2}{\overline{X_2}} \ge \frac{c_3}{\overline{X_3}} \ge \dots \ge \frac{c_n}{\overline{X_n}}$$

## 3.41

Let D(t) and  $T_i(t)$  be as in the solution of Problem 3.31. The inequality in the hint is evident from Figure 3.30, and therefore it will suffice to show that

We have

$$W = R/(1 - \rho)$$

where

$$R = \lim_{t \to \infty} \left\{ \frac{1}{t} \sum_{i=1}^{M(t)} \frac{1}{2} X_i^2 + \frac{1}{t} \sum_{i=1}^{L(t)} \frac{V_i^2}{2} \right\}$$

where L(t) is the number of vacations (or busy periods) up to time t. The average length of an idle period is

 $I = \int_0^{\infty} p(v) \left[ \int_0^V v \lambda e^{-\lambda \tau} d\tau + \int_V^{\infty} \tau \lambda e^{-\lambda \tau} d\tau \right] dv$ 

and it can be seen that the steady-state time average number of vacations per unit time

$$\lim_{t \to \infty} \frac{L(t)}{t} = \frac{1 - \rho}{I}$$

We have

$$\lim_{t \to \infty} \frac{1}{t} \sum_{i=1}^{L(t)} \frac{V_i^2}{2} = \lim_{t \to \infty} \frac{L(t)}{t} \frac{\sum_{i=1}^{L(t)} \frac{V_i^2}{2}}{L(t)} = \lim_{t \to \infty} \frac{L(t)}{t} \frac{\overline{V}^2}{2I} = \frac{\overline{V}^2(1-\rho)}{2I}$$

Therefore

$$R = \frac{\lambda \overline{X}^2}{2} + \frac{\overline{V}^2(1-\rho)}{2I}$$

and

$$W = \frac{\lambda \overline{X}^2}{2(1-\rho)} + \frac{\overline{V}^2}{2I}$$

3.47

(a) Since arrival times and service times are independent, the probability that there was an arrival in a small interval  $\delta$  at time  $\tau$  - x and that this arrival is still being served at time  $\tau$  is  $\lambda \delta [1 - F_X(x)]$ .

(b) We have

$$\overline{X} = \int_{0}^{\infty} x dF_{X}(x)$$

and by calculating the shaded area of the figure below in two different ways we obtain

$$\int x dF_X(x) = \int_0^\infty [1 - F_X(x)] dx$$

This proves the desired expression.





For  $n \ge 1$  we have

$$p_{n}(x - \delta) = \{1 - \lambda [1 - F_{X}(x)]\delta\}p_{n}(x) + \lambda [1 - F_{X}(x)]\delta p_{n-1}(x)$$

and for n = 0 we have

$$p_0(x - \delta) = \{1 - \lambda [1 - F_X(x)]\delta\} p_0(x).$$

Thus  $p_n(x)$ , n = 0, 1, 2, ... are the solution of the differential equations

$$dp_n/dx = a(x)p_n(x) - a(x)p_{n-1}(x)$$
 for  $n \ge 1$ 

$$dp_0/dx = a(x)p_0(x)$$
 for  $n = 0$ 

where

 $\mathbf{a}(\mathbf{x}) = \lambda [1 - \mathbf{F}_{\mathbf{X}}(\mathbf{x})].$ 

Using the known conditions

$$p_n(\infty) = 0 \qquad \text{for } n \ge 1$$
$$p_0(\infty) = 1$$

it can be verified by induction starting with n = 0 that the solution is

$$p_{n}(x) = [e^{x}] \frac{\left[\int_{x}^{\infty} a(y)dy\right]^{n}}{\frac{x}{n!}}, \quad x \ge 0, \ n = 0, \ 1, \ 2, \ ...$$

Since

$$\int_{0}^{\infty} a(y)dy = \lambda \int_{0}^{\infty} [1 - F_X(x)]dy = \lambda E\{X\}$$

we obtain

$$p_n(0) = e^{-\lambda E\{X\}} \frac{[\lambda E\{X\}]^n}{n!}$$
,  $n = 0, 1, 2, ...$ 

Thus the number of arrivals that are still in the system have a steady state Poisson distribution with mean  $\lambda E\{X\}$ .

3.48

(a) Denote

$$f(x) = E_r[(max\{0,r-x\})^2]$$

and

$$g(x) = (E_r[max{0,r-x}])^2,$$

where  $E_r[\cdot]$  denotes expected value with respect to r (x is considered constant). We will prove that f(x)/g(x) is montonically nondecreasing for x nonnegative and thus attain its minimum value for x=0. We have